# The Urban Wage Premium in Historical Perspective

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In the U.S. today, workers that live in an urban area earn pprox 22% more than non-urban workers

• This pattern holds in cities across the world and over (at least) the last 80 years in the US (Boustan, Bunten, and Hearey, 2013)

### Self-selection and the premium

While the raw data shows a large difference in earnings, it is not clear if this is a causal effect or due to non-random sorting

• College-educated and high-skilled workers are increasingly migrating to cities (Autor, 2019; Combes, Duranton, and Gobillon, 2008; Diamond, 2016; Glaeser, Kolko, and Saiz, 2001)

Indeed, panel data estimates shows the current causal wage premium is much smaller

- Glaeser and Maré (2001) estimate effects between 5 to 10%
- For blue-collar workers, Gould (2007) estimate a wage premium of only 1.2 percent

### Why study the historical urban wage premium?

Over time the economic structure of a city has changed

- In the 1940s, cities were primarily manufacturing-based
- Moving towards service-sector in the 21st century (Diamond, 2016)

Understanding the wage premium over the 20th century provides insights into how the benefits of urbanity changes over a country's development

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• Understanding the productivity benefits of cities in manufacture-based economies has strong implications for growth policies in developing countries

### Past approach to estimating causal wage premium

The current literature focuses on a 'movers' approach to estimating the causal wage premium

Comparing individuals who enter or exit an urban location allows to control for 'worker fundamentals' via a fixed-effect strategy

• e.g. Glaeser and Maré (2001), Yankow (2006), Baum-Snow and Pavan (2012), and Card, Rothstein, and Yi (2024)

# Studying the historical urban wage premium is hard

Without panel data, the standard approach is infeasible for estimating a historical causal wage premium

We use the approach of Altonji and Mansfield (2018) to estimate causal urban-wage premium in cross-sectional census data

• Leverage properties of the canonical location-choice model to control for differences in average unobservables between urban and non-urban locations

**Data and Empirical Strategy** 

**Causal Wage Premium** 

**Heterogeneity Results** 

#### Data Sources

We use various U.S. Census surveys for each decade from 1940 to 2010:

- 1940: complete count census
- 1960-1970: 5% census samples
- 1950, 1980, 1990, and 2000: 1% census sample
- 2010: American Community Survey

We restrict the sample to

- adult males between ages 25 and 65,
- and report being employed for at least 26 weeks in the year

#### Observable covariates

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Later, we consider a measure of 'real wage' premium that adjusts for housing costs. Following Ganong and Shoag (2017), we form an individual measure of housing costs

• 5% of reported housing value or 12 $\times$  reported monthly rent. Divide by 48 to get the weekly housing costs

### Definition of Urban Area

Our definition of 'urban' is defined as the Metropolitan Statistical Areas

• MSAs cover urban core and suburbs where residents live

For non-urban areas, we group together all counties not in an MSA by state

• 48 non-urban areas for people to live-in



## **Empirical Strategy**

We consider the following regression specifications:

```
\log(\mathsf{wage}_i) = \beta_0 + \beta_1 \mathsf{Urban}_i + \mathbf{W}_i \delta + u_i
```

Regressions are run separately by year and  $\hat{\beta}_1$  is our wage-premium estimate for that year

•  $W_i$  are a set of worker characteristics that vary across specification

### Individual level Controls

Without covariates, the raw premium (no covariates) is what we presented on the first slide

As a first pass, we control for a set of individual characteristics

- 5-year age bins
- Years of schooling
- indicator for being White

#### Controlling for Individual-level Characteristics



#### Individual level Controls

Takeaway: The individual controls absorb a small portion of the raw premium

• Much less explained than suggested by the standard "urban-living college worker" story

The estimates in later periods are much larger compared to panel data estimates

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# Controlling for Unobservables

In this section, we lay out the methodology to estimate the causal urban wage premium in our cross-sectional setting

• Follow closely to Altonji and Mansfield (2018)

Say individuals sort based on unobserved characteristics  $X_i^U$ . The difference in log wages is conflated by differences in average  $X^U$  in urban and non-urban locations

 $\implies$  It would be sufficient to control for location-specific  $ar{Z}_{\ell(i)}$ , though these are unobservable

#### Location-choice model

The canonical location-choice model has worker i choosing over locations,  $\ell$ , to maximize their indirect utility

$$V_{i\ell} = \boldsymbol{W}_i \boldsymbol{A}_\ell - P_\ell + \varepsilon_{i\ell}$$

- $oldsymbol{A}_\ell$  is a vector of location-specific amenities
- $P_\ell$  is the housing costs in location  $\ell$
- $\varepsilon_{i\ell}$  are idiosyncratic preference draws (assumed Type-1 EV)

To allow for rich sorting patterns, the willingness to pay for amenities varies by worker,  $oldsymbol{W}_i$ 

• E.g. college workers have preferences for certain consumptive amenities (Diamond, 2016)

# Sorting by Characteristics

Since locations vary by their amenities  $A_\ell$ , the kinds of individuals they attract will vary too

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• In equilibrium, the location-average of observable characteristics are a function of the location's amenities  $\bar{X}_\ell = f(A_\ell)$ 

While we can not observe them, the location-average of *unobservable characteristics* are also functions of amenities  $\bar{X}^U_\ell = f^U(A_\ell)$ 

• The mapping and even which amenities impact sorting can differ for observable and unobservable characteristics.

$$ar{m{X}}_\ell = f(m{A}_\ell)$$
 and  $ar{m{X}}^U_\ell = f^U(m{A}_\ell)$ 

Say we knew the mappings f and  $f^U$ . Under conditions we will list below, we can 'invert' observable averages to learn about amenities

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$$\boldsymbol{A}_{\ell} = f^{-1}(\bar{\boldsymbol{X}}_{\ell})$$

Then apply  $f^U$  to learn about unobservables

$$\bar{\boldsymbol{X}}_{\ell}^{U} = f^{U}(f^{-1}(\bar{\boldsymbol{X}}_{\ell}))$$

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 $\implies$  Upshot: Controlling *linearly* for  $\bar{X}_{\ell}$  in our wage premium model is sufficient to control for differences in unobservables between urban and non-urban locations

#### **Necessary Assumptions**

The first three assumptions are standard:

A1 The indirect utility function is  $V_{i\ell} = W_i A_\ell - P_\ell + \varepsilon_{i\ell}$ 

- A2  $\varepsilon_{i\ell}$  are iid type-1 Extreme Value
  - $\rightarrow \,$  standard in urban/trade literature

A3 When choosing locations, workers take prices and amenities as given

ightarrow They do not internalize their own effect on prices and amenities

#### **Necessary Assumptions**

The next two assumptions restrict the relationship between observables, unobservables, and amenities. Remember our mapping relies on:

Observables teach us about amenities

A4 The function  $A_{\ell} = f(\bar{X}_{\ell})$  is invertible

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Amenities teach us about unobservables

A5 The set of location-specific amenities that drive sorting on observables contains all amenities that drive sorting based on unobservable covariates

### Result from Altonji and Mansfield (2018)

With assumptions A1-A5, the location choice model implies the mapping between  $\bar{X}_{\ell}$  and  $\bar{X}_{\ell}^{U}$  is linear

We leverage this to isolate the *causal* urban wage premium by modifying our estimation equation to include a set of location-averages of observable characteristics

$$\log(\mathsf{wage}_i) = \beta_0 + \beta_1 \mathsf{Urban}_i + W_i \delta + \bar{X}_\ell \gamma + u_i$$

### Group Averages

For our group-averages, we include variables we think explain differences in WTP for amenities:

- the share of white workers (the share of non-white workers drops out)
- share of residents in each 5-year age bin
- share of veterans
- share of married and together, married and apart, and single
- share with < HS, HS degree, some college, BA degree, and post-secondary education
- a measure of market access calculated as in Jaworski and Kitchens (2019)

# Plausibility of A4

While there are theoretically many amenities people can 'shop with their feet for', many are tightly bundled together

• Using a principle components analysis on a wide array of amenities, Diamond (2016) finds 5 'bundles' of amenities across locations

Therefore, we need the effective rank of our group averages to be 5 or larger

• Running a similar PCA, we find 6 components explain more than 5 percent of the variance across years

# Plausibility of A5

A5 is a more difficult assumption to assess since it is an assumption about unobservable characteristics

• Later, we think our ability to match estimates in 2000 and 2010 help alleviate concerns about failing to control for unobservables

#### Causal Urban Wage Premium Estimates



# Takeaways

From 1990–2010, the causal premium estimate is around 5%, in line with results from Glaeser and Maré (2001) and Gould (2007)

• This gives us confidence in our ability to control for sorting based on unobservables

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Looking into past decades, the causal estimates grow over time, peaking between 12 and 15% in the middle of the 20th century

In line with the historical evidence:

- In middle 20th century, manufacturing bases in cities offering higher wages
- Decline in wage premium concurrent with sub-urbanization

### Wages net of Housing Costs

Of course, urban areas also face higher housing costs

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We rerun our analysis using  $\log(\mathsf{wage}_i - P_\ell)$ 

- Following Ganong and Shoag (2017), 5% of reported housing value or 12× reported monthly rent. Divide by 48 to get the weekly housing costs
  - ightarrow Averaged over individuals to abstract from individual preferences for housing
  - ightarrow The 1950 census had no housing estimates, so we use an aggregate housing cost measure from the Census

#### Raw Urban (Net of Housing) Wage Premium Estimates



#### Causal Urban (Net of Housing) Wage Premium Estimates



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## Heterogeneity by City Size

Next, we explored heterogeneity in the causal wage premium.

First, the literature has shown that larger cities have a larger causal premium (Baum-Snow and Pavan, 2012)

• We interact the urban indicator with being a top-20 city in terms of population (based on 1940 population)

# Heterogeneity by City Size



Indeed we see the wage premium is moderately larger across years (Baum-Snow and Pavan, 2012), but the difference has shrunk in recent decades

• Consistent with the theory of sorting being increasingly important

# Heterogeneity by Census Region



- The Midwest had a much higher wage premium, consistent with cities' historical premium being driven by manufacturing
- A recent rise in the West's premium, consistent with technology sector

# Heterogeneity for groups of workers

Next, we present three forms of worker-heterogeneity: college education and race

• For each, we take include separate group-averages by race/education because the sorting patterns (and hence linear mappings) might differ by these categories

### Heterogeneity by College Education



The wage-premium is smaller for non-college educated workers, consistent with (Gould, 2007)

### Heterogeneity by Race



The relatively higher wage premium by Black workers occurs during the period of the Great Migration (Boustan, 2016)

### Regional Heterogeneity for Black Workers



At risk of stretching the data too thin, can see the Midwest and Western premiums served as a 'pull factor' for the Great Migration

### Summary

Estimate a causal urban wage premium from 2010 back until 1940

• The causal premium was between 12 and 15% in 1940–1960 and has shrunk over time to around 5%

Heterogeneity analysis in line with historical narrative:

- High-manufacturing in the Midwest gave high wage premia, particularly for Black workers
- College-educated workers benefit more from urbanity

These result help us understand how agglomeration benefits accrue to workers at different points in economic development

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